

# Data-processing inequality with expectations of random variables

(C'est le nom du lemme-clé de cet exposé)

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## Data-processing inequality with expectations of random variables

Data-processing inequality

$$\text{KL}(\mathbb{P}^X, \mathbb{Q}^X) \leq \text{KL}(\mathbb{P}, \mathbb{Q})$$

$X$  is any random variable  $\Omega \rightarrow E$

Our version

$$\text{KL}\left(\text{Ber}(\mathbb{E}_{\mathbb{P}}[X]), \text{Ber}(\mathbb{E}_{\mathbb{Q}}[X])\right) \leq \text{KL}(\mathbb{P}, \mathbb{Q})$$

for  $X : \Omega \rightarrow [0, 1]$

KL is the Kullback-Leibler divergence between two distributions,

$$\text{KL}(\mathbb{P}, \mathbb{Q}) = \begin{cases} +\infty & \text{if } \mathbb{P} \not\ll \mathbb{Q} \\ \int_{\Omega} \left( \frac{d\mathbb{P}}{d\mathbb{Q}} \ln \frac{d\mathbb{P}}{d\mathbb{Q}} \right) d\mathbb{Q} & \text{if } \mathbb{P} \ll \mathbb{Q} \end{cases}$$

# Proof of $KL(P^X, Q^X) \leq KL(P, Q)$ — part 1/2

Proof: We may assume that  $P \ll Q$ , otherwise  $KL(P, Q) = +\infty$  and the inequality is true. We show that we then have

$$P^X \ll Q^X, \quad \text{with} \quad \frac{dP^X}{dQ^X} = \mathbb{E}_Q \left[ \frac{dP}{dQ} \mid X = \cdot \right] \stackrel{\text{not.}}{=} \gamma$$

ie,  $\gamma(x) = \mathbb{E}_Q \left[ \frac{dP}{dQ} \mid X = x \right]$ .

Indeed, for all  $B \in \mathcal{F}$ :

$$\begin{aligned}
 P^X(B) &= P\{X \in B\} = \int_{\Omega} \mathbb{1}_B(x) \frac{dP}{dQ} dQ \stackrel{\text{tower rule}}{=} \int_{\Omega} \mathbb{1}_B(x) \mathbb{E}_Q \left[ \frac{dP}{dQ} \mid X \right] dQ \\
 &\stackrel{\text{not.}}{=} \int_{\Omega} \mathbb{1}_B(x) \gamma(x) dQ \stackrel{\text{by definition of } Q^X}{=} \int_{\Omega} \mathbb{1}_B \gamma dQ^X.
 \end{aligned}$$

# Proof of $KL(P^X, Q^X) \leq KL(P, Q)$ — part 2/2

Therefore,

$$\begin{aligned}
 KL(P^X, Q^X) &= \int_{\Omega^X} \gamma \ln \gamma \, dQ^X = \int_{\Omega} \gamma(x) \ln \gamma(x) \, dQ \\
 &= \int_{\Omega} \left( E_Q \left[ \frac{dP}{dQ} \mid X \right] \ln E_Q \left[ \frac{dP}{dQ} \mid X \right] \right) dQ \quad \begin{array}{l} \text{definition} \\ \text{of} \\ \gamma \end{array} \\
 &\leq \int_{\Omega} E_Q \left[ \frac{dP}{dQ} \ln \frac{dP}{dQ} \mid X \right] dQ \quad \begin{array}{l} \text{conditional} \\ \text{version of} \\ \text{Jensen's inequality} \end{array} \\
 &\stackrel{\text{tower rule}}{=} \int_{\Omega} \left( \frac{dP}{dQ} \ln \frac{dP}{dQ} \right) dQ = KL(P, Q)
 \end{aligned}$$

Reference: Ali and Silvey [1966]; implies joint convexity of KL

Proof of  $KL(\text{Ber}(\mathbb{E}_{\mathbb{P}}[X]), \text{Ber}(\mathbb{E}_{\mathbb{Q}}[X])) \leq KL(\mathbb{P}, \mathbb{Q})$

$$\begin{aligned}
 KL(\underbrace{(\mathbb{P}\alpha\eta)^{\mathbb{1}_E}}_{\text{Ber}(\mathbb{P}\alpha\eta(E))}, \underbrace{(\mathbb{Q}\alpha\eta)^{\mathbb{1}_E}}_{\text{Ber}(\mathbb{Q}\alpha\eta(E))}) &\leq KL(\mathbb{P}\alpha\eta, \mathbb{Q}\alpha\eta) \\
 &= KL(\mathbb{P}, \mathbb{Q}) + KL(\eta, \eta) \\
 &\quad \uparrow \text{if product distributions} \\
 &= KL(\mathbb{P}, \mathbb{Q})
 \end{aligned}$$

Thus:  $KL(\text{Ber}(\mathbb{P}\alpha\eta(E)), \text{Ber}(\mathbb{Q}\alpha\eta(E))) \leq KL(\mathbb{P}, \mathbb{Q})$

The proof is concluded by picking  $E \in \mathcal{F} \otimes \mathcal{B}([0,1])$  such that  $\mathbb{P}\alpha\eta(E) = \mathbb{E}_{\mathbb{P}}[X]$  and  $\mathbb{Q}\alpha\eta(E) = \mathbb{E}_{\mathbb{Q}}[X]$

Namely,  $E = \{(w, x) \in \Omega \times [0,1] : x \leq X(w)\}$

By Tonelli's theorem:

$$\begin{aligned}
 \mathbb{P}\alpha\eta(E) &= \int_{\Omega} \left( \int_{[0,1]} \mathbb{1}_{\{x \leq X(w)\}} d\eta(x) \right) d\mathbb{P}(w) \\
 &= \int_{\Omega} X(w) d\mathbb{P}(w) = \mathbb{E}_{\mathbb{P}}[X]
 \end{aligned}$$

and a similar equality for  $\mathbb{Q}\alpha\eta(E)$ .

## First application: $K$ -armed bandits

$K$  probability distributions  $\nu_1, \dots, \nu_K$

with expectations  $\mu_1, \dots, \mu_K$

$$\longrightarrow \mu^* = \max_{k=1, \dots, K} \mu_k$$

At each round  $t = 1, 2, \dots$ ,

1. Statistician picks **arm**  $I_t \in \{1, \dots, K\}$ , possibly using  $U_{t-1}$
2. She gets a reward  $Y_t$  with law  $\nu_{I_t}$  given  $I_t$
3. This is the **only feedback** she receives

→ **Exploration–exploitation dilemma**

estimate the  $\nu_k$  **vs.** get high rewards  $Y_t$

**Regret:**

$$R_T = \sum_{t=1}^T (\mu^* - \mathbb{E}[Y_t]) = \sum_{k=1}^K \left( (\mu^* - \mu_k) \mathbb{E} \left[ \sum_{t=1}^T \mathbb{I}_{\{I_t=k\}} \right] \right)$$

Indeed,  $Y_t | I_t \sim \nu_{I_t}$ , thus  $\mathbb{E}[Y_t | I_t] = \mu_{I_t}$

Summary:

At each round, pick  $I_t$  (based on  $U_{t-1}$  + past) and get  $Y_t \mid I_t \sim \nu_{I_t}$

Control the regret  $R_T = \sum_{k=1}^K (\mu^* - \mu_k) \mathbb{E}[N_k(T)]$ , where  $N_k(T) = \sum_{t=1}^T \mathbb{I}_{\{I_t=k\}}$

Lower bound  $R_T \iff$  Lower bound  $\mathbb{E}[N_k(T)]$  for  $\mu_k < \mu^*$

Randomized strategy  $\psi = (\psi_t)_{t \geq 0}$ : measurable functions

$$\psi_t : H_t = (U_0, Y_1, U_1, \dots, Y_t, U_t) \mapsto \psi_t(H_t) = I_{t+1}$$

Take  $U_0, U_1, \dots$  iid  $\sim \mathcal{U}_{[0,1]}$  and denote by  $\mathfrak{m}$  the Lebesgue measure

Transition kernel (conditional distributions):

$$\mathbb{P}(Y_{t+1} \in B, U_{t+1} \in B' \mid H_t) = \nu_{\psi_t(H_t)}(B) \mathfrak{m}(B')$$

Summary: history  $H_t = (U_0, Y_1, U_1, \dots, Y_t, U_t)$  and  $I_{t+1} = \psi_t(H_t)$

Lower bound  $\mathbb{E}[N_k(T)]$  for  $\mu_k < \mu^*$ , where  $N_k(T) = \sum \mathbb{I}_{\{I_t=k\}}$

Transition kernel:  $\mathbb{P}(Y_{t+1} \in B, U_{t+1} \in B' \mid H_t) = \nu_{\psi_t(H_t)}(B) m(B')$

**Change of measure:**  $\underline{\nu} = (\nu_1, \dots, \nu_K)$  vs.  $\underline{\nu}' = (\nu'_1, \dots, \nu'_K)$

**Fundamental inequality:** quantifies the impact of such a change

For all  $Z$  taking values in  $[0, 1]$  and  $\sigma(H_T)$ -measurable,

$$\begin{aligned} \sum_{k=1}^K \mathbb{E}_{\underline{\nu}}[N_k(T)] \text{KL}(\nu_k, \nu'_k) &= \text{KL}(\mathbb{P}_{\underline{\nu}}^{H_T}, \mathbb{P}_{\underline{\nu}'}^{H_T}) \\ &\geq \text{KL}\left(\text{Ber}(\mathbb{E}_{\underline{\nu}}[Z]), \text{Ber}(\mathbb{E}_{\underline{\nu}'}[Z])\right) \end{aligned}$$

Inequality holds by the data-processing inequality with expectations of r.v.

**Later use:**  $\underline{\nu}'$  only differ from  $\underline{\nu}$  at  $k$  and  $Z = N_k(T)/T$

# Proof of the equality: chain rule for KL

$$H_{t+1} = (H_t, (Y_{t+1}, U_{t+1})) \text{ and } \mathbb{P}(Y_{t+1} \in B, U_{t+1} \in B' \mid H_t) = \nu_{\psi_t(H_t)}(B) \mathfrak{m}(B')$$

$$\begin{aligned} & \text{KL}(\mathbb{P}_{\underline{\nu}}^{H_{t+1}}, \mathbb{P}_{\underline{\nu}'}^{H_{t+1}}) \\ &= \text{KL}(\mathbb{P}_{\underline{\nu}}^{H_t}, \mathbb{P}_{\underline{\nu}'}^{H_t}) + \text{KL}(\mathbb{P}_{\underline{\nu}}^{(Y_{t+1}, U_{t+1}) \mid H_t}, \mathbb{P}_{\underline{\nu}'}^{(Y_{t+1}, U_{t+1}) \mid H_t}) \\ &= \text{KL}(\mathbb{P}_{\underline{\nu}}^{H_t}, \mathbb{P}_{\underline{\nu}'}^{H_t}) + \mathbb{E}_{\underline{\nu}} \left[ \mathbb{E}_{\underline{\nu}'} \left[ \text{KL}(\nu_{\psi_t(H_t)} \otimes \mathfrak{m}, \nu'_{\psi_t(H_t)} \otimes \mathfrak{m}) \mid H_t \right] \right] \\ &= \text{KL}(\mathbb{P}_{\underline{\nu}}^{H_t}, \mathbb{P}_{\underline{\nu}'}^{H_t}) + \mathbb{E}_{\underline{\nu}} \left[ \mathbb{E}_{\underline{\nu}'} \left[ \text{KL}(\nu_{\psi_t(H_t)}, \nu'_{\psi_t(H_t)}) \mid H_t \right] \right] \\ &= \text{KL}(\mathbb{P}_{\underline{\nu}}^{H_t}, \mathbb{P}_{\underline{\nu}'}^{H_t}) + \mathbb{E}_{\underline{\nu}} \left[ \sum_{k=1}^K \text{KL}(\nu_k, \nu'_k) \mathbb{I}_{\{I_{t+1}=k\}} \right] \end{aligned}$$

By induction: 
$$\text{KL}(\mathbb{P}_{\underline{\nu}}^{H_T}, \mathbb{P}_{\underline{\nu}'}^{H_T}) = \sum_{k=1}^K \mathbb{E}_{\underline{\nu}}[N_k(T)] \text{KL}(\nu_k, \nu'_k)$$

References: already present in Auer, Cesa-Bianchi, Freund and Schapire [2002]

**Fundamental inequality:** For all  $Z \in [0, 1]$  and  $\sigma(H_T)$ -measurable,

$$\sum_{k=1}^K \mathbb{E}_{\underline{\nu}}[N_k(T)] \text{KL}(\nu_k, \nu'_k) \geq \text{kl}(\mathbb{E}_{\underline{\nu}}[Z], \mathbb{E}_{\underline{\nu}'}[Z])$$

**How to use it?**

Bandit problem  $\underline{\nu} = (\nu_1, \dots, \nu_K)$  where  $k$  is suboptimal:  $\mu_k < \mu^*$

Pick  $Z = N_k(T)/T$

Pick  $\underline{\nu}'$  that only differs from  $\underline{\nu}$  at  $k$ :

$$\underline{\nu}' = (\nu_1, \dots, \nu_{k-1}, \nu'_k, \nu_{k+1}, \dots, \nu_K)$$

**Then**

$$\mathbb{E}_{\underline{\nu}}[N_k(T)] \text{KL}(\nu_k, \nu'_k) \geq \text{KL}\left(\text{Ber}(\mathbb{E}_{\underline{\nu}}[N_k(T)/T]), \text{Ber}(\mathbb{E}_{\underline{\nu}'}[N_k(T)/T])\right)$$

To lower bound  $R_T = \sum_{k=1}^K (\mu^* - \mu_k) \mathbb{E}[N_k(T)]$ , lower bound  $\mathbb{E}[N_k(T)]$  for  $\mu_k < \mu^*$

Bandit model  $\mathcal{D}$ : where the  $\nu_1, \dots, \nu_K$  may lie in

**Assumption (UFC – uniform fast convergence on  $\mathcal{D}$ )**

*The strategy  $\psi$  is such that:*

*For all bandit problems  $\underline{\nu} = (\nu_1, \dots, \nu_K)$  in  $\mathcal{D}$ , for all  $\mu_k < \mu^*$ ,*

$$\forall \alpha \in (0, 1], \quad \mathbb{E}_{\underline{\nu}}[N_k(T)] = o(T^\alpha)$$

Bandit problem  $\underline{\nu} = (\nu_1, \dots, \nu_K)$  where  $k$  is suboptimal:  $\mu_k < \mu^*$

Pick  $\nu'_k \in \mathcal{D}$  with expectation  $\mu'_k > \mu^*$

Form  $\underline{\nu}'$  that only differs from  $\underline{\nu}$  at  $k$ :

$$\underline{\nu}' = (\nu_1, \dots, \nu_{k-1}, \nu'_k, \nu_{k+1}, \dots, \nu_K)$$

**Then**  $\mathbb{E}_{\underline{\nu}}[N_k(T)] = o(T)$  and  $T - \mathbb{E}_{\underline{\nu}'}[N_k(T)] = o(T^\alpha)$

$$\mathbb{E}_{\underline{\nu}}[N_k(T)] \text{KL}(\nu_k, \nu'_k) \geq \text{KL}\left(\text{Ber}(\mathbb{E}_{\underline{\nu}}[N_k(T)/T]), \text{Ber}(\mathbb{E}_{\underline{\nu}'}[N_k(T)/T])\right)$$

$$\mathbb{E}_{\underline{\nu}}[N_k(T)] = o(T) \text{ and } T - \mathbb{E}_{\underline{\nu}'}[N_k(T)] = o(T^\alpha) \text{ for a strategy } \psi \text{ UFC on } \mathcal{D}$$

$$\text{Also, } \text{KL}(\text{Ber}(p), \text{Ber}(q)) = p \ln \frac{p}{q} + (1-p) \ln \frac{1-p}{1-q} \geq (1-p) \ln \frac{1}{1-q} - \ln 2$$

Combining all these elements:

$$\begin{aligned} & \mathbb{E}_{\underline{\nu}}[N_k(T)] \\ & \geq \frac{1}{\text{KL}(\nu_k, \nu'_k)} \text{kl}\left(\mathbb{E}_{\underline{\nu}}[N_k(T)/T], \mathbb{E}_{\underline{\nu}'}[N_k(T)/T]\right) \\ & \geq \frac{1}{\text{KL}(\nu_k, \nu'_k)} \left( -\ln 2 + \left(1 - \mathbb{E}_{\underline{\nu}}[N_k(T)/T]\right) \ln \frac{1}{1 - \mathbb{E}_{\underline{\nu}'}[N_k(T)/T]} \right) \\ & \geq \frac{1}{\text{KL}(\nu_k, \nu'_k)} \left( -\ln 2 + (1 - o(1)) \ln \frac{1}{T^{\alpha-1}} \right) \end{aligned}$$

$$\text{Thus, } \forall \alpha \in (0, 1], \quad \liminf_{T \rightarrow \infty} \frac{\mathbb{E}_{\underline{\nu}}[N_k(T)]}{\ln T} \geq \frac{1}{\text{KL}(\nu_k, \nu'_k)} \frac{\ln T^{1-\alpha}}{\ln T}$$

That is, for **all models  $\mathcal{D}$**  (for the first time, **no assumption on  $\mathcal{D}$** )  
for all strategies  $\psi$  **UFC** on  $\mathcal{D}$  (this is not a real restriction)  
for all bandit problems  $\underline{\nu} = (\nu_1, \dots, \nu_K)$  in  $\mathcal{D}$   
for  $\mu_k < \mu^*$

### Lemma

for all  $\nu'_k$  in  $\mathcal{D}$  with  $\mu'_k > \mu^*$ ,

$$\liminf_{T \rightarrow \infty} \frac{\mathbb{E}_{\underline{\nu}}[N_k(T)]}{\ln T} \geq \frac{1}{\text{KL}(\nu_k, \nu'_k)}$$

**Theorem** (see Lai and Robbins [1985], Burnetas and Katehakis [1996])

$$\liminf_{T \rightarrow \infty} \frac{\mathbb{E}_{\underline{\nu}}[N_k(T)]}{\ln T} \geq \frac{1}{\mathcal{K}_{\text{inf}}(\nu_k, \mu^*, \mathcal{D})}$$

where  $\mathcal{K}_{\text{inf}}(\nu_k, \mu^*, \mathcal{D}) = \inf \{ \text{KL}(\nu_k, \nu'_k) : \nu'_k \in \mathcal{D} \text{ with } \mu'_k > \mu^* \}$

This distribution-dependent bound is **asymptotically optimal**:

$$\liminf_{T \rightarrow \infty} \frac{\mathbb{E}_{\underline{\nu}}[N_k(T)]}{\ln T} \geq \frac{1}{\mathcal{K}_{\text{inf}}(\nu_k, \mu^*, \mathcal{D})}$$

I.e., at least for well-behaved models  $\mathcal{D}$ , we can exhibit a matching upper bound:

$$\limsup_{T \rightarrow \infty} \frac{\mathbb{E}_{\underline{\nu}}[N_k(T)]}{\ln T} \leq \frac{1}{\mathcal{K}_{\text{inf}}(\nu_k, \mu^*, \mathcal{D})}$$

See Lai and Robbins [1985], Burnetas and Katehakis [1996], Honda and Takemura [2010–2015], Cappé, Garivier, Maillard, Munos and Stoltz [2013], etc.

Replacing the  $o(T^\alpha)$  in the definition of UFC by a  $O(\ln T)$ :

$$\mathbb{E}_{\underline{\nu}}[N_k(T)] \geq \frac{\ln T}{\mathcal{K}_{\text{inf}}(\nu_k, \mu^*, \mathcal{D})} - \Omega(\ln(\ln T))$$

Cf. the upper bound of Honda and Takemura [2015]:

This **second-order term**  $-\ln(\ln T)$  is **optimal**

## Second application: Fano's inequality

For collections of distributions  $\mathbb{P}_\theta$ ,  $\mathbb{Q}_\theta$  and r.v.  $Z_\theta$ :

$$\begin{aligned} & \text{KL}\left(\text{Ber}\left(\int_{\Theta} \mathbb{E}_{\mathbb{P}_\theta}[Z_\theta] d\nu(\theta)\right), \text{Ber}\left(\int_{\Theta} \mathbb{E}_{\mathbb{Q}_\theta}[Z_\theta] d\nu(\theta)\right)\right) \\ & \leq \int_{\Theta} \text{KL}\left(\text{Ber}\left(\mathbb{E}_{\mathbb{P}_\theta}[Z_\theta]\right), \text{Ber}\left(\mathbb{E}_{\mathbb{Q}_\theta}[Z_\theta]\right)\right) d\nu(\theta) \\ & \leq \int_{\Theta} \text{KL}(\mathbb{P}_\theta, \mathbb{Q}_\theta) d\nu(\theta) \end{aligned}$$

Thus,

$$\int_{\Theta} \mathbb{E}_{\mathbb{P}_\theta}[Z_\theta] d\nu(\theta) \leq \frac{\int_{\Theta} \text{KL}(\mathbb{P}_\theta, \mathbb{P}_0) d\nu(\theta) + \ln 2}{-\ln \int_{\Theta} \mathbb{E}_{\mathbb{P}_0}[Z_\theta] d\nu(\theta)}$$

In particular,

$$\frac{1}{N} \sum_{i=1}^N \mathbb{P}_i(A_i) \leq \frac{\frac{1}{N} \sum_{i=1}^N \text{KL}(\mathbb{P}_i, \mathbb{P}_0) + \ln 2}{-\ln \frac{1}{N} \sum_{i=1}^N \mathbb{P}_0(A_i)}$$

We obtained, with no real assumptions:

$$\int_{\Theta} \mathbb{E}_{\mathbb{P}_{\theta}}[Z_{\theta}] d\nu(\theta) \leq \frac{\int_{\Theta} \text{KL}(\mathbb{P}_{\theta}, \mathbb{P}_0) d\nu(\theta)}{-\ln \int_{\Theta} \mathbb{E}_{\mathbb{P}_0}[Z_{\theta}] d\nu(\theta)}$$

To get lower bounds on the posterior concentration rate, we use

$$Z_{\theta} = \mathbb{P}_{\pi}(\theta' : \|\theta' - \theta\|_2 \leq \varepsilon_n \mid \mathbf{X}_{1:n})$$

E.g., when  $\mathbb{P}_{\theta} = \mathcal{N}(\theta, \sigma^2)^{\otimes n}$ :

$$\inf_{\theta \in \Theta} \mathbb{E}_{\theta} \left[ \mathbb{P}_{\pi}(\theta' : \|\theta' - \theta\|_2 \leq \varepsilon_n \mid \mathbf{X}_{1:n}) \right] \leq \frac{n\rho^2 \varepsilon_n^2 / (2\sigma^2) + \ln 2}{d \ln \rho} \leq c < 1$$

for  $\varepsilon_n$  of order  $\sigma \sqrt{d/n}$

We took  $\nu$  as the uniform distribution over  $B(0, \rho \varepsilon_n)$