

We first complete the proof of the distribution-dependent lower bound.

Protocol: with randomized strategies $(\Psi_t)_{t \geq 0}$, where

$$\Psi_t: H_t = (U_0, y_{t,1}, U_1, \dots, y_{t,K}, U_t) \mapsto \Psi_t(H_t) = I_{t+1}$$

We have:
$$\begin{cases} y_t | I_t \sim \mathcal{Y}_{I_t} & \forall t \\ U_0, U_1, \dots \text{ iid} \sim \mathcal{U}_{[0,1]} \end{cases}$$

Regret:
$$R_T = T\mu^* - \mathbb{E} \left[\sum_{t=1}^T y_t \right] \stackrel{\text{by the tower rule}}{=} \sum_{a=1}^K \Delta_a \mathbb{E}_y [N_a(T)]$$

where $\mu_a = \mathbb{E}(y_a)$, $\mu^* = \max_{a=1, \dots, K} \mu_a$,
 $\Delta_a = \mu^* - \mu_a$

and \mathbb{E}_y denotes the expectation when the underlying bandit problem is $\mathcal{Y} = (y_a)_{a \in \{1, \dots, K\}}$

↳ To control the regret (upper or lower bound it), it is thus necessary and sufficient to control the $\mathbb{E}_y [N_a(T)]$.

Definitions:

- Bandit model \mathcal{D} : A subset of the set $\mathcal{M}_1(\mathbb{R})$ of all probability distributions over $(\mathbb{R}, \mathcal{B}(\mathbb{R}))$ with an expectation.

The decision-maker knows \mathcal{D} but does not know the specific bandit problem $\mathcal{Y} = (y_a)_{a \in \{1, \dots, K\}} \in \mathcal{D}^K$ at hand.

- Strategy Ψ consistent w.r.t. a model \mathcal{D} : when $\forall \mathcal{Y} \in \mathcal{D}^K, \forall \epsilon \in (0, 1], \forall n \text{ s.t. } n \geq \frac{1}{\epsilon}$
 $\mathbb{E}_y [N_a(T)] = o(T^\epsilon)$

Lemma (Fundamental inequality for stochastic bandits): For all strategies Ψ ,

For all bandit problems $\mathcal{Y} = (y_a)_{a \in \{1, \dots, K\}}$ and $\mathcal{Y}' = (y'_a)_{a \in \{1, \dots, K\}} \in \mathcal{D}^K$
 with $y'_a \leq y_a$ for all a ,

For all random variables Z taking values in $[0, 1]$ and that are $\sigma(H_T)$ -measurable,

$$\begin{aligned} \sum_{a=1}^K \mathbb{E}_y [N_a(T)] \text{KL}(y_a, y'_a) &= \text{KL}(\mathbb{P}_{\mathcal{Y}}^{H_T}, \mathbb{P}_{\mathcal{Y}'}^{H_T}) \\ &\geq \text{KL}(\text{Ber}(\mathbb{E}_y[Z]), \text{Ber}(\mathbb{E}_{\mathcal{Y}'}[Z])) \end{aligned}$$

where $\mathbb{P}_{\mathcal{Y}}^{H_T}$ and $\mathbb{P}_{\mathcal{Y}'}^{H_T}$ denote the laws of H_T when the strategy is Ψ and when the underlying bandit problems are respectively \mathcal{Y} and \mathcal{Y}' .

Proof: • The inequality \geq is a direct application of the data-processing inequality with expectations, see the previous lecture for its statement.

• For the equality: We will explain how $\mathbb{P}_{\mathcal{Y}}^{H_T}$ is constructed.

With no loss of generality, we can consider that

- the underlying probability space is $\Omega = \{0,1\} \times (\mathbb{R} \times \{0,1\})^T$

- H_T is the identity over Ω , i.e., that the $U_0, Y_1, U_1, \dots, Y_T, U_T$ are the projections on each component,

- $\mathbb{P}_{\mathcal{Y}}$ is given by

$$\forall B \in \mathcal{B}(\{0,1\}), \quad \mathbb{P}_{\mathcal{Y}}(U_0 \in B) = \eta(B)$$

$$\forall t \in \{0, \dots, T-1\}, \quad \forall B' \in \mathcal{B}(\mathbb{R}), \quad \forall B \in \mathcal{B}(\{0,1\}), \quad \mathbb{P}_{\mathcal{Y}}(Y_{t+1} \in B' \text{ and } U_{t+1} \in B \mid H_t) = \int_{\mathcal{Y}_t(H_t)} \nu(B) \eta(B)$$

where $\mathcal{B}(S)$ is the Borel- σ -algebra of a set $S \subseteq \mathbb{R}$
 η is the Lebesgue measure over $\{0,1\}$

In particular: $\mathbb{P}_{\mathcal{Y}}^{H_t}$ refers to the first $1+t$ marginals of $\mathbb{P}_{\mathcal{Y}}^{H_T} = \mathbb{P}_{\mathcal{Y}}$.

A similar construction can be done for the bandit problem \mathcal{Y}' .

Now, the equality $\mathbb{P}_{\mathcal{Y}}(Y_{t+1} \in B' \text{ and } U_{t+1} \in B \mid H_t) = \int_{\mathcal{Y}_t(H_t)} \nu(B) \eta(B)$

indicates that $\mathbb{P}_{\mathcal{Y}}^{H_{t+1}} = K_t \mathbb{P}_{\mathcal{Y}}^{H_t}$ for the regular transition kernel

$$K_t(h, \cdot) = \int_{\mathcal{Y}_t(h)} \nu(\cdot) \ll \eta$$

regularity is:
 for $E \in \mathcal{B}(\mathbb{R}) \cap \mathcal{B}(\{0,1\})$
 $h \mapsto K_t(h, E)$ is measurable

Similarly, $\mathbb{P}_{\mathcal{Y}'}^{H_{t+1}} = K'_t \mathbb{P}_{\mathcal{Y}'}^{H_t}$ for $K'_t(h, \cdot) = \int_{\mathcal{Y}'_t(h)} \nu'(\cdot) \ll \eta$

Let us check the assumptions of our chain rule:

(*) $\forall h, \quad K_t(h, \cdot) \ll K'_t(h, \cdot)$ as $\forall a, \nu'_a \ll \nu_a$ by assumption

(**) $(h, (y, u)) \mapsto \frac{dK_t(h, \cdot)}{dK'_t(h, \cdot)}(y, u) = \sum_{a=1}^K \mathbb{1}_{\int \nu'_t(h) = a} \frac{d\nu'_t}{d\nu_t}(y)$
 is indeed bi-measurable.

Therefore, for $t \in \{0, \dots, T-1\}$,

$$\begin{aligned} \text{KL}(\mathbb{P}_{\mathcal{Y}}^{H_{t+1}}, \mathbb{P}_{\mathcal{Y}'}^{H_{t+1}}) &= \text{KL}(\mathbb{P}_{\mathcal{Y}}^{H_t}, \mathbb{P}_{\mathcal{Y}'}^{H_t}) + \int \text{KL}(\nu_{\int \nu'_t(h)} \ll \eta, \nu'_{\int \nu'_t(h)} \ll \eta) d\mathbb{P}_{\mathcal{Y}}^{H_t}(h) \\ &= \text{KL}(\mathbb{P}_{\mathcal{Y}}^{H_t}, \mathbb{P}_{\mathcal{Y}'}^{H_t}) + \sum_{a=1}^K \text{KL}(\nu_a, \nu'_a) \mathbb{P}_{\mathcal{Y}}^{H_t} \left\{ \int \nu'_t(h) = a \right\} \end{aligned}$$

Now, $I_{\mathcal{F}_{t+1}} = \Psi_t(H_t)$ so that

$$\mathbb{P}_y^{H_t} \{ \Psi_t(H_t) = a_j \} = \mathbb{P}_y \{ \Psi_t(H_t) = a_j \} = \mathbb{P}_y \{ I_{\mathcal{F}_{t+1}} = a_j \} \\ = \mathbb{E}_y [\mathbb{1}_{\{I_{\mathcal{F}_{t+1}} = a_j\}}]$$

Summing up:

$$- \text{KL}(\mathbb{P}_y^{H_0}, \mathbb{P}_y^{H_0}) = \text{KL}(\mathbb{P}_y^{U_0}, \mathbb{P}_y^{U_0}) = \text{KL}(\eta, \eta) = 0$$

$$- \forall t \in \{0, \dots, T-1\}, \quad \text{KL}(\mathbb{P}_y^{H_{t+1}}, \mathbb{P}_y^{H_{t+1}}) = \text{KL}(\mathbb{P}_y^{H_t}, \mathbb{P}_y^{H_t}) \\ + \sum_{a=1}^k \text{KL}(\nu_{a,1}^y, \nu_a^y) \mathbb{E}_y [\mathbb{1}_{\{I_{\mathcal{F}_{t+1}} = a\}}]$$

so that the stated result follows by induction.

Back to the theorem giving the lower bound:

Theorem:

For all bandit models $\mathcal{D} \subset \mathcal{M}_1(\mathbb{R})$,

For all (possibly randomized) strategies Ψ consistent w.r.t \mathcal{D} ,

For all bandit problems $\nu = (\nu_a)_{a \in \{1, \dots, k\}} \in \mathcal{D}^k$,

For all suboptimal arms a (ie, arms a with $\Delta_a > 0$),

$$\liminf_{T \rightarrow +\infty} \frac{\mathbb{E}_\Psi [N_a(T)]}{\ln T} \geq \frac{1}{\text{KL}(\nu_a, \mu^*, \mathcal{D})}$$

where $\text{KL}(\nu_a, \mu^*, \mathcal{D}) = \inf \{ \text{KL}(\nu_a, \nu_a') : \nu_a' \in \mathcal{D} \text{ with } \mathbb{E}(\nu_a') > \mu^* \}$
with the convention $\inf \emptyset = +\infty$.

Correction for Exercise #3:

The proof of this theorem based on the lemma above was your Exercise #2.

See next page.

Proof: We have $K_{\text{inf}}(\nu_a, \mu^*) = \inf \{ \text{KL}(\nu_a, \nu_a^i) : \nu_a^i \in \mathcal{D} \text{ and } \mathbb{E}(\nu_a^i) \geq \mu^* \}$
 $= \inf \{ \text{KL}(\nu_a, \nu_a^i) : \nu_a^i \in \mathcal{D}, \nu_a \ll \nu_a^i \text{ and } \mathbb{E}(\nu_a^i) \geq \mu^* \}$
 (cf. convention: $\inf \emptyset = +\infty$ and the fact that $\text{KL}(\nu_a, \nu_a^i) = +\infty$ when $\nu_a \not\ll \nu_a^i$)

This is why we will

- Fix $\mathcal{D}, \Psi, \bar{\nu}$ and a s.t. $\Delta_a > 0$
- Fix an alternative model ν^j of the form

$$\begin{cases} \nu_k^j = \bar{\nu}_k & \forall k \neq a \\ \nu_a^j & \text{s.t. } \nu_a^j \in \mathcal{D}, \nu_a^j \ll \bar{\nu}_a \text{ and } \mathbb{E}(\nu_a^j) \geq \mu^* \end{cases}$$

That is, $\bar{\nu}$ and ν^j only differ at a ; a is the unique optimal arm in ν^j

- Take $Z = N_a(T)/T$ which is indeed $[0,1]$ -valued $\sigma(H_T)$ -measurable

Our fundamental inequality yields, since $\bar{\nu}$ and ν^j only differ at a :

$$\mathbb{E}_{\bar{\nu}}[N_a(T)] \text{KL}(\bar{\nu}_a, \nu_a^j) \geq \text{KL}(\text{Ber}(\mathbb{E}_{\bar{\nu}}[N_a(T)/T]), \text{Ber}(\mathbb{E}_{\nu^j}[N_a(T)/T])) \\ \geq -\ln 2 + (1 - \mathbb{E}_{\nu^j}[N_a(T)/T]) \ln \frac{1}{1 - \mathbb{E}_{\bar{\nu}}[N_a(T)/T]}$$

indeed: $\text{KL}(\text{Ber}(p), \text{Ber}(q))$

$$= p \ln \frac{p}{q} + (1-p) \ln \frac{1-p}{1-q}$$

$$= \underbrace{p \ln \frac{1}{q}}_{\geq 0} + (1-p) \ln \frac{1}{1-q} + \underbrace{(p \ln p + (1-p) \ln(1-p))}_{\geq -\ln 2} \\ \geq -\ln 2 + (1-p) \ln \frac{1}{1-q}$$

for all $p, q \in (0,1)$ and even for all $p, q \in [0,1]$ (study the cases $q=0$ and $q=1$ separately)

Now, the considered strategy Ψ is consistent and:

- in the problem $\bar{\nu}$, a is suboptimal: $\mathbb{E}_{\bar{\nu}}[N_a(T)/T] \rightarrow 0$

- in the problem \mathcal{V}^j , all arms $k \neq a$ are suboptimal:

$$\text{for all } \alpha \in (0, 1], \quad T - \mathbb{E}_{\mathcal{V}^j} [N_a(T)] = \sum_{k \neq a} \mathbb{E}_{\mathcal{V}^j} [N_k(T)] = o(T^\alpha)$$

↳ in particular, for T large enough,

$$\frac{1}{1 - \mathbb{E}_{\mathcal{V}^j} [N_k(T)/T]} = \frac{T}{T - \mathbb{E}_{\mathcal{V}^j} [N_k(T)]} \geq \frac{T}{T^\alpha} = T^{1-\alpha}$$

Substituting back and dividing by $\ln T$: for all $\alpha \in (0, 1]$, for T large enough:

$$\frac{\mathbb{E}_{\mathcal{V}^j} [N_k(T)]}{\ln T} \text{KL}(\mathcal{V}_{a_1}^j, \mathcal{V}_{a_2}^j) \geq -\frac{\ln 2}{\ln T} + \underbrace{\left(1 - \mathbb{E}_{\mathcal{V}^j} \left[\frac{N_k(T)}{T}\right]\right)}_{\rightarrow 0} \frac{\ln T^{1-\alpha}}{\ln T} = 1 - \alpha$$

thus

$$\liminf_{T \rightarrow +\infty} \frac{\mathbb{E}_{\mathcal{V}^j} [N_k(T)]}{\ln T} \text{KL}(\mathcal{V}_{a_1}^j, \mathcal{V}_{a_2}^j) \geq 1 - \alpha$$

Letting $\alpha \rightarrow 0$:

$$\liminf_{T \rightarrow +\infty} \frac{\mathbb{E}_{\mathcal{V}^j} [N_k(T)]}{\ln T} \text{KL}(\mathcal{V}_{a_1}^j, \mathcal{V}_{a_2}^j) \geq 1$$

Whether $\text{KL}(\mathcal{V}_{a_1}^j, \mathcal{V}_{a_2}^j) < +\infty$ or $= +\infty$, we thus get

$$\liminf_{T \rightarrow +\infty} \frac{\mathbb{E}_{\mathcal{V}^j} [N_k(T)]}{\ln T} \geq \frac{1}{\text{KL}(\mathcal{V}_{a_1}^j, \mathcal{V}_{a_2}^j)}$$

The left-hand side is independent of $\mathcal{V}_{a_2}^j \in \mathcal{D}$ s.t. $\mathcal{V}_{a_2}^j \succcurlyeq \mathcal{V}_{a_1}^j$ and $E(\mathcal{V}_{a_1}^j) \succcurlyeq \mu^*$, so that taking the supremum of the right-hand side over these $\mathcal{V}_{a_2}^j$, we get the desired $\frac{1}{\text{KL}(\mathcal{V}_{a_1}^j, \mu^*)}$ lower bound.

Remark: Non-asymptotic lower bounds are possible (for a super-consistent \rightarrow strategies and a well-behaved \rightarrow models), but are heavily technical.

Exercise 1. $\frac{1}{\text{Kinf}(\mathcal{X}_T, \mu^*, \mathcal{D})}$ vs. $\frac{8}{\Delta_n^2}$ for UCB

Recall that in the model $\mathcal{D} = \mathcal{J}([0,1])$, the UCB algorithm enjoys the following performance bound:

$$\forall \mathcal{J} \in \mathcal{J}([0,1])^K, \quad \forall a \text{ s.t. } \Delta_n > 0, \\ \mathbb{E}_\mathcal{J} [N_k(T)] \leq \frac{8}{\Delta_n^2} \ln T + 2.$$

Actually, there are refinements of UCB that get the distribution-dependent constant $\frac{8}{\Delta_n^2}$ arbitrarily close to $\frac{2}{\Delta_n^2}$.

But how do these $\frac{8}{\Delta_n^2}$ and $\frac{2}{\Delta_n^2}$ constants compare to $\frac{1}{\text{Kinf}(\mathcal{X}_T, \mu^*, \mathcal{J}([0,1]))}$?

(1) For $p, q \in [0,1]$, we denote

$$k(p, q) = \text{KL}(\text{Ber}(p), \text{Ber}(q))$$

Shows that $\forall (p, q) \in [0,1]^2, \quad k(p, q) \geq 2(p-q)^2$.

(2) Show Pinsker's inequality: let (Ω, \mathcal{F}) be a measurable space, let \mathbb{P}, \mathbb{Q} be two distributions over (Ω, \mathcal{F}) , then:

$$\| \mathbb{P} - \mathbb{Q} \|_{\text{TV}} = \sup_{A \in \mathcal{F}} | \mathbb{P}(A) - \mathbb{Q}(A) | \leq \sqrt{\frac{1}{2} \text{KL}(\mathbb{P}, \mathbb{Q})}$$

↑
the total variation distance between \mathbb{P} and \mathbb{Q}

Even better, show the stronger form: $\sup_{Z: \mathcal{F}\text{-measurable taking values in } [0,1]} | \mathbb{E}_\mathbb{P}[Z] - \mathbb{E}_\mathbb{Q}[Z] | \leq \sqrt{\frac{1}{2} \text{KL}(\mathbb{P}, \mathbb{Q})}$

(3) Exhibit a lower bound on $\text{Kinf}(\mathcal{X}_T, \mu^*, \mathcal{J}([0,1]))$ and conclude that some work is needed to get an upper bound matching our lower bound!

Lower bound on the regret for adversarial bandits
 \leftrightarrow Minimax lower bound on the regret for stochastic bandits

We discussed so far $\mathbb{E}_\nu [N_a(T)]$ in terms of distribution-dependent bounds like $C_a(\nu) \ln T$

What about distribution-free / minimax bounds?

To that end we restrict our attention to the model $\mathcal{D} = \mathcal{P}([0,1])$, the set of all probability distributions over $[0,1]$.

Stochastic bandits

With each arm a is associated $\nu_a \in \mathcal{P}([0,1])$

For $t=1,2,\dots$

- The decision maker picks $I_t \in \{1..K\}$
- Her reward Y_t , which is such that $Y_t | I_t \sim \nu_{I_t}$, is her only piece of information

Aim: control the regret

$$\bar{R}_T = T \max_{a=1..K} \mathbb{E}(\nu_a) - \mathbb{E} \left[\sum_{t=1}^T Y_t \right]$$

Adversarial bandits

An opponent selects the payoffs g_{jt}

For $t=1,2,\dots$

- The opponent picks $(g_{jt} - g_{kt}) \in [0,1]^K$ while, simultaneously,
- The decision-maker picks $I_t \in \{1..K\}$
- Her payoff is $g_{I_t t}$ and this is the only piece of information she gets

Aim: control the regret

$$R_T = \max_{k=1..K} \sum_{t=1}^T g_{kt} - \sum_{t=1}^T g_{I_t t}$$

Typical adversarial results

(Auer, Cesa-Bianchi, Freund, Schapire, 2002, later improved by Audibert and Bubeck, 2009):

Strategies

such that for all opponents picking gains in $[0,1]$, for all $T \geq 1$,

$$\left\{ \begin{array}{l} \text{with probability at least } 1-\delta, \\ \mathbb{E}[R_T] \leq C \sqrt{TK \ln(K/\delta)} \end{array} \right. \quad R_T \leq C \sqrt{TK \ln(K/\delta)}$$

where

the probability and \mathbb{E} are w.r.t decision-maker's internal randomization

for some numerical constant C

For "oblivious" opponents (ie, when the g_{jt} do not "react" to the decision-maker's actions):

the $\sqrt{\ln K}$ can be dropped.

It is in particular the case when $g_{jt} \sim \nu_a$ i.e. in an independent way

In this stochastic model:

$$\begin{aligned} \mathbb{E}[R_T] &= \mathbb{E}\left[\max_{k=1..K} \sum_{t=1}^T g_{kt}\right] - \mathbb{E}\left[\sum_{t=1}^T g_{I_t,t}\right] \\ &\geq T \max_{k=1..K} \mathbb{E}[g_{k,1}] - \mathbb{E}\left[\sum_{t=1}^T y_t\right] \\ &= T \max_{k=1..K} \mathbb{E}(y_k) - \mathbb{E}\left[\sum_{t=1}^T y_t\right] = \bar{R}_T \end{aligned}$$

$g_{I_t,t}$ is y_t

The adversarial results entail in particular that there exists a strategy of the decision-maker such that

$$\bar{R}_T \leq \sup_{\gamma_1, \dots, \gamma_K \in \mathcal{P}([0,1])} \mathbb{E}[R_T] \leq C \sqrt{TK}$$

for some numerical constant C

Lower bound:

Theorem: For all (randomized) strategies of the decision-maker, for all $K \geq 2$ and $T \geq K \ln 2 / 2$,

$$\mathbb{E}[R_T] \geq \sup_{\substack{\text{opponents} \\ \text{picking } g_{jt} \in [0,1]}} \mathbb{E}[R_T] \geq \sup_{\gamma_1, \dots, \gamma_K \in \mathcal{P}([0,1])} \bar{R}_T \geq \frac{1}{20} \sqrt{TK}$$

the only inequality we need to prove!

Proof: (i) Bandit problem $\gamma^{(0)} = (\text{Ber}(\frac{1}{2}), \dots, \text{Ber}(\frac{1}{2}))$

versus bandit problems $\gamma^{(k)} = (\text{Ber}(\frac{1}{2}), \dots, \text{Ber}(\frac{1}{2}), \text{Ber}(\frac{1}{2} + \varepsilon), \text{Ber}(\frac{1}{2}), \dots)$
 for $i \in \{1, \dots, K\}$ and $\varepsilon \in (0, \frac{1}{2})$
 in i -th position

There exists $k \in \{1, \dots, K\}$ such that $\mathbb{E}_{\gamma^{(0)}} [N_k(T)] \leq T/K$. For this k :

The fundamental inequality indicates that

$$\begin{aligned} \sum_{a=1}^K \mathbb{E}_{\gamma^{(0)}} [N_a(T)] \text{KL}(\gamma_a^{(0)}, \gamma_a^{(k)}) &\geq \text{kl}\left(\mathbb{E}_{\gamma^{(0)}} \left[\frac{N_k(T)}{T}\right], \mathbb{E}_{\gamma^{(k)}} \left[\frac{N_k(T)}{T}\right]\right) \\ &= \mathbb{E}_{\gamma^{(0)}} [N_k(T)] \text{KL}(\gamma_k^{(0)}, \gamma_k^{(k)}) \leq \frac{T}{K} \text{kl}\left(\frac{1}{2}, \frac{1}{2} + \varepsilon\right) \end{aligned}$$

Using Pinsker's inequality:

$$\begin{aligned} \text{kl}\left(\mathbb{E}_{\gamma^{(0)}} \left[\frac{N_k(T)}{T}\right], \mathbb{E}_{\gamma^{(k)}} \left[\frac{N_k(T)}{T}\right]\right) \\ \geq 2 \left(\mathbb{E}_{\gamma^{(0)}} \left[\frac{N_k(T)}{T}\right] - \mathbb{E}_{\gamma^{(k)}} \left[\frac{N_k(T)}{T}\right]\right)^2 \end{aligned}$$

Solving for $\mathbb{E}_{\mathcal{Y}^{(k)}} \left[\frac{N_k(T)}{T} \right]$:

$$\begin{aligned} \mathbb{E}_{\mathcal{Y}^{(k)}} \left[\frac{N_k(T)}{T} \right] &\leq \mathbb{E}_{\mathcal{Y}^{(0)}} \left[\frac{N_k(T)}{T} \right] + \sqrt{\frac{T}{2K} \text{kl}\left(\frac{1}{2}, \frac{1}{2} + \varepsilon\right)} \\ &\leq \frac{1}{K} + \sqrt{\frac{T}{2K} \text{kl}\left(\frac{1}{2}, \frac{1}{2} + \varepsilon\right)} \\ &\leq \frac{1}{K} + \sqrt{\frac{(2 \ln 2) T \varepsilon^2}{K}} \quad \text{for } \varepsilon \in (0, \frac{1}{2\sqrt{2}}] \end{aligned}$$

Indeed: $\varepsilon \in (0, \frac{1}{2\sqrt{2}}]$ entails $4\varepsilon^2 \leq \frac{1}{2}$, as $\ln\left(\frac{1}{1-u}\right) \leq (2 \ln 2)u \quad \forall u \in (0, \frac{1}{2}]$
and $\text{kl}\left(\frac{1}{2}, \frac{1}{2} + \varepsilon\right) = \frac{1}{2} \ln\left(\frac{1}{1-2\varepsilon}\right) + \frac{1}{2} \ln\left(\frac{1}{1+2\varepsilon}\right) = \frac{1}{2} \ln \frac{1}{1-4\varepsilon^2} \leq (4 \ln 2) \varepsilon^2$

(2) Let's go back to the regret.

$$\begin{aligned} \text{In } \mathcal{Y}^{(k)}: \quad \bar{R}_T &= \sum_{a \neq k} \varepsilon \mathbb{E}_{\mathcal{Y}^{(k)}} [N_a(T)] = \varepsilon (T - \mathbb{E}_{\mathcal{Y}^{(k)}} [N_k(T)]) \\ &\quad \begin{array}{c} \uparrow \\ \text{gap} \\ \text{of } a \text{ from } a \\ \text{number of} \\ \text{times } a \text{ is pulled} \end{array} \\ &\geq T \varepsilon \left(\underbrace{1 - \frac{1}{K}}_{\geq \frac{1}{2}} - \sqrt{\frac{(2 \ln 2) T \varepsilon^2}{K}} \right) \end{aligned}$$

to be optimized over $\varepsilon \in (0, \frac{1}{2\sqrt{2}}]$

$$\text{Optimal } \varepsilon \text{ s.t. } \frac{1}{2} - 2\varepsilon \sqrt{\frac{(2 \ln 2) T}{K}} = 0, \quad \text{ie, } \varepsilon = \frac{1}{4} \times \frac{1}{\sqrt{\frac{(2 \ln 2) T}{K}}}$$

which is $\leq \frac{1}{2\sqrt{2}}$ as soon as $T \ln 2 / K \geq \frac{1}{2}$,

For this ε , the lower bound is

ie $T \geq K \ln 2 / 2$

$$T \varepsilon \left(\frac{1}{2} - \varepsilon \sqrt{\frac{(2 \ln 2) T}{K}} \right) = \frac{T \varepsilon}{4} = \frac{1}{16 \sqrt{2 \ln 2}} \sqrt{TK} \geq \frac{1}{20} \sqrt{TK}$$

Note about the constants:

improve the $\frac{1}{20} \sqrt{TK}$ bound by further restricting the considered T (ie, by imposing $T \geq \gamma K$ with a larger γ), we can get $c' \sqrt{TK}$ with $c' > \frac{1}{20}$.

This is done in the upper bound on $\ln\left(\frac{1}{1-u}\right)$: eg, $\leq 4 \ln(4/3) u$ for $u \in [0, 1/4]$

$$\hookrightarrow \text{Entails: } \sup \bar{R}_T \geq \frac{\sqrt{2}-1}{\sqrt{32 \ln(4/3)}} \sqrt{TK} \geq 0.136 \sqrt{TK} \quad \text{for } T \geq \frac{1}{4 \ln(4/3)} K \approx 0.87 K$$

Open question:

So, the minimax rates for the regret for stochastic bandits or oblivious opponents are \sqrt{TK} .

What is the minimax rate against general, reactive opponents?

→ Should the upper bound be improved?

→ Should the proof technique be improved?
(in particular, look for sequences of payoffs with strong correlations/dependencies in \mathcal{I}_t with the past).

Stochastic bandits :What about arms indexed by a continuum?

Setting 1 : Arms indexed by $x \in A$, where A is some possibly large set
 With each arm $x \in A$ is associated a probability distribution ν_x over \mathbb{R} s.t. $E(\nu_x)$ exists
 At each round, the decision-maker picks $I_t \in A$, gets a reward Y_t drawn at random according to ν_{I_t} (given I_t); and this is the only feedback she gets.

Definition : $f: x \in A \mapsto E(\nu_x)$ is the mean-payoff function

Regret :

$$\bar{R}_T = T \sup_{x \in A} f(x) - E \left[\sum_{t=1}^T Y_t \right]$$

Setting 2 : [special case] \longrightarrow Noisy optimization of a function.

We fix $f: A \rightarrow \mathbb{R}$

The noise is given by a sequence of iid random variables

$\varepsilon_1, \varepsilon_2, \dots$

When $I_t \in A$ is picked, $Y_t = f(I_t) + \varepsilon_t$

\hookrightarrow Special case of setting #1 where ν_x is the distribution of $f(x) + \varepsilon_1$ (all these distributions have the same shape given by the common distribution of the ε_j)

We of course need conditions for the regret to be minimized.

Definition : Let \mathcal{F} be a set of possible bandit problems $\mathcal{F} = (\nu_x)_{x \in A}$

The regret can be controlled (in a non-uniform way) against \mathcal{F} if :

we also say that (A, \mathcal{F}) is tractable

there exists a strategy

s.t. $\forall \mathcal{F} \in \mathcal{F}, \bar{R}_T = o(T)$.

Ex: $A = \{1, \dots, K\}$ and $\mathcal{F} = (\mathcal{P}([0,1]))^K$, the set of all K -tuples of probability distributions over $[0,1]$

the case of finitely many arms with bounded distributions

→ UCB does the job.

Counter-example: $A = [0,1]$ and $\mathcal{F} = (\mathcal{P}([0,1]))^{[0,1]}$

↑ illustrating that continuity is a minimal requirement.

all bandit problems $(\nu_x)_{x \in [0,1]}$ with distributions ν_x having support $[0,1]$.

In fact: Consider $(\delta_0)_{x \in [0,1]}$ the bandit problem in which each arm x is associated with the Dirac mass on 0

Fix any strategy: it gets $Y_t = 0$ $\forall t$ and uses a sequence of (possibly) random choices $I_t, t \geq 1$

Since probability distributions can only have at most countably many atoms,

$$\mathcal{Y} = \{x \in [0,1] : \exists t \mid \mathbb{P}\{I_t = x\} > 0 \text{ under } (\delta_0)_{x \in [0,1]}\}$$

is countable. In particular, $[0,1] \setminus \mathcal{Y}$ is non empty.

But the strategy behaves the same under the problem $(\nu_x)_{x \in [0,1]}$ in which

$$\begin{cases} \nu_x^1 = \delta_0 & \forall x \neq x_0 \\ \nu_{x_0}^1 = \delta_1 & \text{for one fixed } x_0 \in [0,1] \setminus \mathcal{Y} \end{cases}$$

With probability 1, it thus never hits x_0 .

Therefore, $Y_t = 0$ a.s. $\forall t$ and $\bar{R}_T = T - \mathbb{E}\left[\sum_{t=1}^T Y_t\right] = T$.

Actually, continuity is necessary and sufficient for the regret to be controlled, as long as A is not too large.

Theorem: Let A be a ^{cont} metric space and let \mathcal{F} be the set of bandit problems $(\nu_x)_{x \in A}$

with: → $\forall x, \nu_x$ is a distribution over $[0,1]$

→ a continuous mean-payoff function $f: x \mapsto \mathbb{E}(\nu_x)$

The regret can be controlled against $\mathcal{F}^{\text{cont}}$ if and only if A is separable.

Corollary. Let \mathcal{F}^{all} be the family of all bandit models $(\mu_x)_{x \in A}$ with distributions ν_x over $[0,1]$. Then the regret against \mathcal{F}^{all} can be controlled if and only if A is at most countable.

Before we prove these facts, consider the following more concrete example, in which, by strengthening the regularity requirement on the mean-payoff function, we can even get rates.

Exercise 2. Let $A = [0,1]$ and let \mathcal{F}^{lip} be the family of bandit models $(\mu_x)_{x \in [0,1]}$ with distributions ν_x over $[0,1]$ and with mean-payoff functions that are Lipschitz.

Exhibit a strategy based on UCB + a sequence of discretizations of $[0,1]$ into K bins (to be refined over time) such that:

Hint:

First, prove a performance bound by splitting $[0,1]$ into $[\frac{i-1}{K}, \frac{i}{K}]$ with $i=1, \dots, K$

for a fixed K , where each bin $[\frac{i-1}{K}, \frac{i}{K}]$ plays the role of an i in a bandit problem with finitely many arms. Then discuss how to pick K over time, as we do in the next proof.

$\forall \mu \in \mathcal{F}^{\text{lip}}$,

$$\bar{R}_T \leq (3L + 6\sqrt{8\ln T} + 2)T^{2/3} + 2$$

where L is the Lipschitz constant of the mean-payoff function of μ .

Proof of the Corollary:

We endow A with the discrete topology, i.e., choose the distance $d(x,y) = \mathbb{1}_{x \neq y}$. Then:

1. All applications $f: A \rightarrow \mathbb{R}$ are continuous
2. A is separable if and only if A is at most countable.

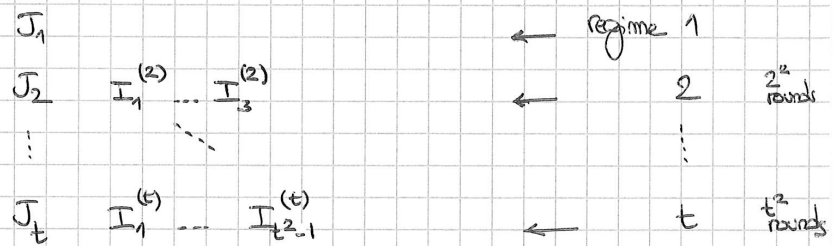
Proof of the Theorem:

It relies on the possibility or impossibility of uniform exploration of the arms.

1) If A is separable: let $(x_n)_{n \geq 1}$ be a collection of points in A that is dense

In particular, the probability distribution $\mu = \sum_{n \geq 1} \frac{1}{2^{nn}} \delta_{x_n}$ is such that $\mu(V) > 0$ for all open sets $V \subset A$.

We pick elements $J_1, J_2, I_1^{(2)}, \dots, J_t, I_1^{(t)}, \dots, I_{t-1}^{(t)}$ as follows:



where $\left\{ \begin{array}{l} \text{the } J_s \text{ are drawn at random according to } \mu \\ \text{the } I_s^{(r)}, 1 \leq s \leq r^2-1, \text{ follow from the UCS strategy with arms } J_1, \dots, J_r \end{array} \right.$

In regime r :

$$r^2 \max_{s \leq r} \mu_{J_s} - \mathbb{E} \left[\sum_{s=1}^{r^2} Y_{S_{r^2+s}} \right]$$

$\leq \underbrace{1}_{\text{for } J_r} + \underbrace{c}_{\text{well-chosen numerical constant}} \sqrt{\underbrace{r^3}_{\text{distribution-free regret bound for UCS on } r^2-1 \text{ steps with } r \text{ arms}} \ln r}$
(we show this bound as an exercise)

regime r starts at time $S_{r-1} + 1$ where $S_i = \sum_{j=1}^i (j-1)^2$

Let $\epsilon > 0$, let \tilde{r}_ϵ the first (random) time when $\mu_{J_r} = f(J_r) \geq \sup_{z \in A} f(z) - \epsilon$

We have $\tilde{r}_\epsilon < +\infty$ a.s. because:

- by continuity of f , there exists an open set V_ϵ with $\forall x \in V_\epsilon, f(x) \geq \sup_A f - \epsilon$;

- we have $\tilde{r}_\varepsilon \leq \inf \{ r \geq 1 : \mathbb{P}_r \in \mathcal{V}_\varepsilon \} < +\infty$ a.s.
 as this random variable follows a geometric distribution with parameter $\rho(\mathcal{V}_\varepsilon) > 0$.

for $r \geq \tilde{r}_\varepsilon$, $\max_{s \leq r} \mu_{J_s} + \varepsilon \geq \sup_A f$

so that
$$\bar{R}_T = T \sup_A f - \mathbb{E} \left[\sum_{t=1}^T J_t \right]$$

$$\leq \sum_{r=1}^{\tilde{r}_\varepsilon - 1} r^2 + T\varepsilon + \sum_{r=\tilde{r}_\varepsilon}^{r_T - 1} (1 + c\sqrt{r^3 \ln r}) + r_T^2$$

↑ lengths of regimes $r \leq \tilde{r}_\varepsilon - 1 < +\infty$ a.s. ↑ regime r_T may be incomplete

the "complete regimes" where r_T is such that time T belongs to regime r_T :
 we have r_T^3 of order T ie r_T of order $T^{1/3}$

and

$$\sum_{r \leq r_T} (1 + c\sqrt{r^3 \ln r}) \leq \sum_{r \leq r_T} (1 + cr^{3/2} \sqrt{\ln r}) = O(r_T^{5/2} \sqrt{\ln r_T}) = O(T^{5/6} \sqrt{\ln T})$$

thus, $\limsup_{T \rightarrow +\infty} \frac{\bar{R}_T}{T} \leq \varepsilon$ a.s.

but since \bar{R}_T is a deterministic quantity and this is true $\forall \varepsilon > 0$, we have

$$\lim_{T \rightarrow +\infty} \frac{\bar{R}_T}{T} = 0 \text{ as requested}$$

2) If A is not separable:

* We use the following characterization of separability (which relies on Zorn's lemma):

A metric space X is separable if and only if it contains no uncountable subset \mathcal{D} s.t.

$$\rho = \inf \{ d(x, y) : x, y \in \mathcal{D} \} > 0.$$

In particular, if A is not separable, there exist an uncountable subset $\mathcal{D} \subset A$ and $\rho > 0$ such that the balls $B(a, \rho/2)$, with $a \in \mathcal{D}$, are all disjoint.

\Rightarrow No probability distribution over A can give a positive mass to all these balls.

* We consider the bandit models $\nu^{(a)}$ inducing mean-payoff functions $f^{(a)} : x \in A \mapsto (1 - \frac{d(x, a)}{\rho/2})^+$; in particular, $\nu^{(a)}_x = \delta_x$ for $x \in B(a, \rho/2)$.
 \uparrow $f^{(a)}$ is indeed continuous.

We proceed as in the example showing the necessity of continuity and consider the bandit model $(\delta_x)_{x \in A}$, as well as any strategy and the laws induced by the \mathcal{I}_t under this model: let ν_t be the law of \mathcal{I}_t under $(\delta_x)_{x \in A}$ and let $\alpha = \sum_{t \geq 1} \frac{1}{2^t} \nu_t$.

There exists $a \in A$ s.t. $\alpha(B(a, \rho/2)) = 0$, that is, s.t., $\forall t \geq 1$, $\mathbb{P}(\mathcal{I}_t \in B(a, \rho/2) \text{ under } (\delta_x)_{x \in A}) = 0$.

The considered strategy is therefore such that the \mathcal{I}_t have the same distribution under $(\delta_x)_{x \in A}$ and $\nu^{(a)}$. In particular,

$\mathbb{E}[\sum_{t=1}^T Y_t] = 0$ in both cases, but in the latter case,

$\sup f^{(a)} = 1$, so that $\bar{R}_T = T$ against $\nu^{(a)}$. The regret is not controlled against $\nu^{(a)} \in \bar{\mathcal{F}}^{\text{cont}}$.