

Le premier exercice de l'examen vous fera prouver le résultat suivant :

Distribution-free (ie uniform) lower bounds.

We prove: For all $K \geq 2$ and $T \geq K/5$

$$\inf_{\text{strategies } \psi} \sup_{\substack{x_1, \dots, x_K \\ \text{in } \mathcal{P}(\{1, J\})}} R_T \geq \frac{1}{20} \sqrt{TK}.$$

and even:

$\sup$ over $x_1, \dots, x_K$ being Bernoulli distributions

The MOSS strategy (Minimax Optimal Strategy in the Stochastic case of bandit problems)

Index policy relying on

$$U_a(t) = \hat{\mu}_a(t) + \sqrt{\frac{1}{2N_a(t)} \ln_+ \left(\frac{t}{K N_a(t)} \right)}$$

for $t \geq K$,

and where $\ln_+ = \max\{\ln, 0\}$

That is:

For $t = 1, \dots, K$: pull arm $A_t = t$

For $t \geq K+1$: pull arm $A_t \in \arg \max_{a=1 \dots K} U_a(t-1)$

Difference to UCB: we replace the exploration bonus $\sqrt{2 \ln t / N_a(t)}$ by $\sqrt{\ln_+ (t / (K N_a(t))) / (2 N_a(t))}$

↳ no exploration after a
was pulled sufficiently
often (t/K times)

We prove a distribution-free bound:

Theorem: MOSS is such that $\sup_{\substack{\mu_1, \dots, \mu_K \\ \text{distributions} \\ \text{over } [0,1]}} \bar{R}_T \leq K-1 + 45 \sqrt{KT}$

(the constant 45 can be improved)

but indeed
minimax optimal as
its name indicates!

Proof.First step: $U_{a^*}(t-1) \leq U_{A_t}(t-1)$ by definition of A_t as an argmax

$$\text{thus } R_T = \sum_{t=1}^T \mathbb{E}[\mu^* - \mu_{A_t}]$$

$$\leq \underbrace{k-1}_{\substack{\text{if we played} \\ \text{each arm once in the} \\ \text{first } k \text{ steps, and} \\ \text{at most } k-1 \text{ were suboptimal}}} + \sum_{t=k+1}^T \mathbb{E}[\mu^* - U_{a^*}(t)] + \sum_{t=k+1}^T \mathbb{E}[U_{A_t}(t) - \mu_{A_t}]$$

$$\leq \sqrt{kT} + \sum_{t=k+1}^T \mathbb{E}[(U_{A_t}(t) - \mu_{A_t})^+ - \sqrt{kT}]$$

Second step:
Control of each

$$\mathbb{E}[\mu^* - U_{a^*}(t)] \text{ term by } 20\sqrt{kT}$$

we write
for $t \geq k$

$$\mathbb{E}[\mu^* - U_{a^*}(t)] \leq \mathbb{E}[(\mu^* - U_{a^*}(t))^+]$$

$$\leq \sum_{l=0}^{+\infty} \mathbb{E}[(\mu^* - U_{a^*}(t))^+ \mathbb{1}_{\substack{N_{a^*}(t) \in [x_{l+1}, x_l] \\ \text{where } x_l = \beta^{-l} t/K \\ \text{for some fixed } \beta > 1 \\ \text{and } l = 0, 1, 2, \dots}}}]$$

$$+ \mathbb{E}[(\mu^* - U_{a^*}(t))^+ \mathbb{1}_{\{N_{a^*}(t) \geq t/k\}}]$$

denoted
by ε_l

$$\text{Now, } U_{a^*}(t) = \mu_{a^*}(t) + \begin{cases} 0 & \text{if } N_{a^*}(t) \geq t/k \\ \sqrt{\frac{1}{2N_{a^*}(t)} \ln \left(\frac{t}{K N_{a^*}(t)} \right)} & \text{if } N_{a^*}(t) \in [x_{l+1}, x_l] \\ \sqrt{\frac{1}{2x_l} \ln \left(\frac{t}{K x_l} \right)} & \text{if } N_{a^*}(t) \geq t/k \end{cases}$$

Therefore,

$$(*) \quad \mathbb{E}[\mu^* - U_{a^*}(t)] \leq \mathbb{E}[(\mu^* - \hat{\mu}_{a^*}(t))^+ \mathbb{1}_{\{N_{a^*}(t) \geq t/k\}}] + \sum_{l=0}^{+\infty} \mathbb{E}[(\mu^* - \hat{\mu}_{a^*}(t) - \varepsilon_l)^+ \mathbb{1}_{\{N_{a^*}(t) \in [x_{l+1}, x_l]\}}]$$

$$\text{Lemma: } \mathbb{E}[(\mu^* - \hat{\mu}_{a^*}(t) - \varepsilon)^+ \mathbb{1}_{\{N_{a^*}(t) \geq n_0\}}] \leq \frac{1}{\sqrt{n_0}} e^{-2n_0 \varepsilon^2}$$

Proof of the lemma:

$$\mathbb{E}[(\mu^* - \hat{\mu}_{a^*}(t) - \varepsilon)^+ \mathbb{1}_{\{N_{a^*}(t) \geq n_0\}}]$$

$$= \int_0^{+\infty} \mathbb{P}[\mu^* - \hat{\mu}_{a^*}(t) - \varepsilon \geq u \text{ and } N_{a^*}(t) \geq n_0] du$$

$$= \int_0^{+\infty} \mathbb{P}\{Z_t^* \geq (\varepsilon u) N_t^*(t) \text{ and } N_t^*(t) \geq n_0\} du$$

where $Z_t^* = N_t^*(t) \left(\mu^* - \hat{\mu}_{N_t^*(t)} \right) = \sum_{s=1}^t \left(\mu^* - \frac{y_s}{\delta} \right) \mathbb{1}_{\{N_s^*(t) = s\}}$ is a martingale, and for all $x \in \mathbb{R}$, $S_{x,t} = e^{xZ_t^* - \frac{x^2}{8} N_t^*(t)}$ is a supermartingale.

Thus, by Markov-Chernoff, we continue the bounding as, for $x > 0$:

$$= \int_0^{+\infty} \mathbb{P}\{e^{xZ_t^* - \frac{x^2}{8} N_t^*(t)} \geq \exp\left(N_t^*(t) \left(x(\varepsilon u) - \frac{x^2}{8}\right)\right) \text{ and } N_t^*(t) \geq n_0\} du$$

$$\leq \int_0^{+\infty} \sum_{\ell=n_0}^{+\infty} e^{-2\ell(\varepsilon u)^2} \mathbb{E}\left[S_{\frac{x(\varepsilon u)}{t}, t} \mathbb{1}_{\{N_t^*(t) = \ell\}}\right] du$$

we pick $x = 4(\varepsilon u)$ so that $x(\varepsilon u) - \frac{x^2}{8} = 2(\varepsilon u)^2$

⊗ independent of ℓ , which will be useful in other proofs!

$$\leq e^{-2n_0\varepsilon^2} \int_0^{+\infty} e^{-2n_0 u^2} \mathbb{E}\left[S_{\frac{x(\varepsilon u)}{t}, t} \mathbb{1}_{\{N_t^*(t) \geq n_0\}}\right] du$$

where $\mathbb{E}\left[S_{\frac{x(\varepsilon u)}{t}, t}\right] \leq 1$

All in all, $\mathbb{E}\left[(\mu^* - \hat{\mu}_{N_t^*(t)} - \varepsilon)^+ \mathbb{1}_{\{N_t^*(t) \geq n_0\}}\right]$

$$\leq e^{-2n_0\varepsilon^2} \int_0^{+\infty} e^{-2n_0 u^2} du = e^{-2n_0\varepsilon^2} \frac{\sqrt{\pi/8n_0}}{\sqrt{n_0}} \leq e^{-2n_0\varepsilon^2} \frac{\sqrt{\pi/8}}{\sqrt{n_0}} \leq 1$$

integral of a Gaussian density, up to the normalization factor

$$\hookrightarrow \sigma^2 = \frac{n_0}{4} \text{ in } \int_0^{+\infty} \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{u^2}{2\sigma^2}} du = \frac{1}{2}$$

Substituting the lemma in (*):

$$\mathbb{E}[\mu^* U_{a^*}(t)] \leq \sqrt{\frac{K}{t}} + \sum_{l=0}^{+\infty} \underbrace{\frac{1}{\sqrt{x_{a^*}}} e^{-2x_{a^*} \varepsilon_e^2}}_{\frac{1}{\sqrt{x_{a^*}}} \exp\left(-2x_{a^*} \frac{1}{2x_e} \ln\left(\frac{t}{Kx_e}\right)\right)} = \sqrt{\frac{K}{t}} \beta^{(2+\frac{1}{2})} \exp\left(-\frac{2}{\beta} \ln \beta\right)$$

where $\sum_{l=0}^{+\infty} \beta^{l(\frac{1}{2}-\frac{1}{\beta})}$ is $< +\infty$ as soon as $\beta \in (1, 2)$

$$= \sqrt{\frac{K}{t}} \beta^{\frac{5}{2}} \exp(-2 \ln \beta) = \sqrt{\frac{K}{t}} \beta^{\frac{1}{2} + l(\frac{1}{2}-\frac{1}{\beta})}$$

E.g., for $\beta = \frac{3}{2}$,

$$\sum_{l=0}^{+\infty} \left(\frac{3}{2}\right)^{\frac{1}{2} + l(\frac{1}{2}-\frac{2}{3})} = \sqrt{\frac{3}{2}} \sum_{l=0}^{+\infty} \alpha^l = \frac{1}{1-\alpha} \sqrt{\frac{3}{2}} \leq 19$$

where $\alpha = \left(\frac{3}{2}\right)^{\frac{1}{2}-\frac{2}{3}} \in (0, 1)$

All in all: we obtain a $\sqrt{\frac{K}{t}} + 19\sqrt{\frac{K}{t}} = 20\sqrt{\frac{K}{t}}$ bound as claimed.

Third step:

$$\sum_{t=K+1}^T \mathbb{E}[(U_{A_t}(t) - \mu_{A_t} - \sqrt{\frac{K}{t}})^+] \leq 4\sqrt{KT}$$

$$= \sum_{t=K}^{T-1} \mathbb{E}[(U_{A_{t+1}}(t) - \mu_{A_{t+1}} - \sqrt{\frac{K}{t}})^+]$$

We decompose the expectations of interest according to the $\{A_t = a\}$ and $\{N_a(t) = l\}$:

$$\sum_{t=K}^{T-1} \mathbb{E}[(U_{A_{t+1}}(t) - \mu_{A_{t+1}} - \sqrt{\frac{K}{t}})^+] = \sum_{a=1}^K \sum_{l=1}^T \sum_{t=K}^{T-1} \mathbb{E}[(U_a(t) - \mu_a - \sqrt{\frac{K}{t}})^+ \mathbb{1}_{\{A_{t+1}=a\}} \mathbb{1}_{\{N_a(t)=l\}}]$$

We now use $(U_a(t) - \mu_a - \sqrt{\frac{K}{t}})^+ \leq (\hat{\mu}_a(t) - \mu_a - \sqrt{\frac{K}{t}})^+ + \begin{cases} 0 & \text{if } N_a(t) \geq \frac{t}{K} \\ \sqrt{\frac{2}{N_a(t)} \ln\left(\frac{t}{K N_a(t)}\right)} & \text{if } N_a(t) < \frac{t}{K} \end{cases}$

also smaller than $\sqrt{\frac{1}{N_a(t)} \ln\left(\frac{T}{K N_a(t)}\right)}$ if $N_a(t) < \frac{t}{K}$

and get therefore the upper bound

$$\sum_{a=1}^K \sum_{l=1}^T \sum_{t=K}^{T-1} \mathbb{E}[(\hat{\mu}_a(t) - \mu_a - \sqrt{\frac{K}{t}})^+ \mathbb{1}_{\{A_{t+1}=a\}} \mathbb{1}_{\{N_a(t)=l\}}] + \sum_{a=1}^K \sum_{l=1}^{\lceil T/K \rceil} \sqrt{\frac{1}{2l}} \ln\left(\frac{T}{Kl}\right) \mathbb{E}\left[\sum_{t=K}^{T-1} \mathbb{1}_{\{N_a(t)=l\}} \mathbb{1}_{\{A_{t+1}=a\}}\right]$$

We will repeatedly use that
 $\forall a, \ell_j \quad \sum_{t=k}^{T-1} \mathbb{1}_{\{N_k(t) = \ell_j\}} \mathbb{1}_{\{A_{t+1} = a\}} \leq 1$ (ie disjoint union)

cf. $N_k(t)$ increases by 1 whenever a is played

Also,
$$\sum_{\ell=1}^{\lceil T/K \rceil} \sqrt{\frac{1}{2\ell} \ln\left(\frac{T}{K\ell}\right)} \leq \int_0^{\lceil T/K \rceil} \sqrt{\frac{1}{2x} \ln\left(\frac{T}{Kx}\right)} dx$$

change of variable $u = T/(Kx)$

$$\leq \sqrt{\frac{T}{2K}} \int_1^{+\infty} u^{-3/2} \sqrt{\ln u} du$$

by the change of variable $u = e^{v^2}$

$$= \sqrt{\frac{T}{2K}} \int_0^{+\infty} 2v^2 e^{-v^2/2} dv = \sqrt{\pi} \sqrt{\frac{T}{K}}$$

Summarizing what we proved so far: $\leq \frac{\sqrt{\pi}}{\sqrt{2}} \sqrt{\frac{T}{K}}$ for each a

$$\sum_{t=K}^{T-1} \mathbb{E}[(U_{A_{t+1}} - \mu_{A_{t+1}} - \sqrt{K_t})^+] \leq \sqrt{\pi} \sqrt{KT} + \sum_{a=1}^K \sum_{\ell=1}^T \sum_{t=K}^{T-1} \mathbb{E}[(\hat{\mu}_a(t) - \mu_a - \sqrt{K_t})^+ \mathbb{1}_{\{A_{t+1}=a\}} \mathbb{1}_{\{N_k(t)=\ell\}}]$$

We resort again to $Z_{a,t} = N_k(t) (\hat{\mu}_a(t) - \mu_a)$ martingale
 and $S_{x,t}^{(a)} = e^{xZ_{a,t} - \frac{x^2}{8} N_k(t)}$ supermartingale
 where $x = 4(\sqrt{K_T} + u)$

For each a_j

$$\sum_{\ell=1}^T \sum_{t=K}^{T-1} \mathbb{E}[(\hat{\mu}_a(t) - \mu_a - \sqrt{K_T})^+ \mathbb{1}_{\{A_{t+1}=a\}} \mathbb{1}_{\{N_k(t)=\ell\}}]$$

$$\stackrel{n=4(\sqrt{K_T}+u)}{\geq} \sum_{\ell=1}^T \sum_{t=K}^{T-1} \int_0^{+\infty} \mathbb{P}\left[\frac{xZ_{a,t}}{-\frac{x^2}{8} N_k(t)} \geq N_k(t) \left(\frac{x(u+\sqrt{K_T})}{-\frac{x^2}{8}}\right) \mid A_{t+1}=a \mid N_k(t)=\ell\right] du$$

$$\leq \sum_{\ell=1}^T \sum_{t=K}^{T-1} \int_0^{+\infty} \underbrace{e^{-2\ell(u+\sqrt{K_T})^2}}_{\leq e^{-2\ell u^2}} e^{-2\ell K_T} \mathbb{E}\left[\frac{S_{x,t}^{(a)}}{\mathbb{1}_{\{A_{t+1}=a\}} \mathbb{1}_{\{N_k(t)=\ell\}}}\right] du$$

by Hoeffding-Chernoff

the sum over t of these will be ≤ 1

issue: this depends on t ...
 but can be replaced in some cases by $S_{x,0}^{(a)} = 1$

But: remember Doob's maximal inequality for non-negative supermartingales:

$$\mathbb{P}\left\{\sup_{t \geq 0} S_{xt}^{(a)} \geq c\right\} \leq \frac{\mathbb{E}[S_{x0}^{(a)}]}{c} = \frac{1}{c}$$

see also an alternative treatment on the next page

Then,

$$\sum_{l=1}^T \sum_{t=K}^{T-1} \mathbb{E}\left[\left(\hat{\mu}_a(t) - \mu_a - \sqrt{\frac{K}{t}}\right)^+ \mathbb{1}_{\{A_{t-1}=a\}} \mathbb{1}_{\{N_a(t)=l\}}\right]$$

as before! $\rightarrow$

$$= \sum_{l=1}^{+\infty} \sum_{t=K}^{T-1} \int_0^{+\infty} \mathbb{P}\left\{S_{xt}^{(a)} \geq e^{2l(u+\sqrt{K/t})^2} \text{ and } A_{t-1}=a \text{ and } N_a(t)=l\right\} du$$

$$\leq \sum_{l=1}^T \int_0^{+\infty} \sum_{t=K}^{T-1} \mathbb{P}\left\{\left(\sup_{s \geq 0} S_{xs}^{(a)}\right) \geq e^{2l(u+\sqrt{K/t})^2} \text{ and } A_{t-1}=a \text{ and } N_a(t)=l\right\} du$$

of disjoint union!

$$\leq \sum_{l=1}^T \int_0^{+\infty} \mathbb{P}\left\{\left(\sup_{s \geq 0} S_{xs}^{(a)}\right) \geq e^{2l(u+\sqrt{K/t})^2}\right\} du$$

Doob's maximal inequality

$$\leq \sum_{l=1}^T \int_0^{+\infty} e^{-2l(u+\sqrt{K/t})^2} du \leq \sum_{l=1}^{+\infty} \frac{1}{\sqrt{e}} e^{-2lK/t}$$

and same treatment as in the lemma of the first part of the proof

This step is concluded by calculations:

$$\begin{aligned} \sum_{l=1}^T \frac{1}{\sqrt{e}} e^{-2lK/t} &\leq \int_0^T \frac{1}{\sqrt{e}} e^{-2xK/t} dx \\ &= \sqrt{\frac{T}{2K}} \int_0^{+\infty} \frac{e^{-u}}{\sqrt{u}} du = \sqrt{\frac{T}{2K}} \int_0^{+\infty} e^{-u^2} du = \sqrt{\frac{\pi}{2}} \sqrt{\frac{T}{K}} \end{aligned}$$

Final bound is: $\sqrt{\pi} \sqrt{KT} + K \sqrt{\frac{\pi}{2}} \sqrt{\frac{T}{K}} \leq 4\sqrt{KT}$

General conclusion: Final bound given by

$$\begin{aligned} K-1 + \left(\sum_{t=K+1}^T 2\alpha\sqrt{\frac{K}{t}}\right) + \sqrt{KT} + 4\sqrt{KT} &\leq K-1 + 5\sqrt{KT} + 2\alpha \int_0^T \sqrt{\frac{K}{t}} dt \\ &= K-1 + 4.5\sqrt{KT} \end{aligned}$$

Alternative treatment (credits to Enzo Miller) of the end of Step #3:

We were stuck at

$$\sum_{l=1}^T \sum_{t=K}^{T-1} \int_0^{+\infty} e^{-2lu^2} e^{-2lK\tau} \mathbb{E} \left[S_{x,t}^{(a)} \mathbb{1}_{\{A_{t+1}=a\}} \mathbb{1}_{\{N_h(t)=l\}} \right] d\tau$$

$$= \sum_{l=1}^T \int_0^{+\infty} e^{-2lu^2} e^{-2lK\tau} \mathbb{E} \left[\underbrace{\sum_{t=K}^{T-1} S_{x,t}^{(a)} \mathbb{1}_{\{A_{t+1}=a\}} \mathbb{1}_{\{N_h(t)=l\}}}_{= S_{x,T_l}^{(a)}} \right] d\tau$$

where T_l is given by:

$$T_l = \inf \{ t \in \{1, \dots, T\} : A_{t+1} = a \text{ and } N_h(t) = l \}$$

We should get $\mathbb{E}[S_{x,T_l}^{(a)}] \leq \mathbb{E}[S_{x,0}^{(a)}] = 1$
 from the optional stopping theorem (« the same
 d'arrêt de Doob ») provided some
 verifications. (T_l should be a bounded
 stopping time.)

Adversarial bandits.

(Rather stated in terms of losses than rewards!)

Setting:At each round $t=1,2,\dots$

1. The opponent and the decision-maker simultaneously choose $\ell_t = (\ell_{jt})_{j \in \{1, \dots, N\}}$ and $I_t \sim p_t$, where $p_t \in \mathcal{P}(\{1, \dots, N\})$
2. The opponent gets to see p_t and I_t ; the decision-maker only observes $\ell_{I_t t}$ (her own loss).

Regret:

$$R_T = \sum_{t=1}^T \ell_{I_t t} - \min_{j=1, \dots, N} \sum_{t=1}^T \ell_{jt}$$

vs. Pseudo-regret:

$$\bar{R}_T = \mathbb{E} \left[\sum_{t=1}^T \ell_{I_t t} \right] - \min_{j=1, \dots, N} \mathbb{E} \left[\sum_{t=1}^T \ell_{jt} \right]$$

↑ same definition as for stochastic bandits, up to the conversion of losses ℓ_{jt} into rewards $M - \ell_{jt}$ (for a well-chosen bound M)

↑ Why $\mathbb{E}$?
cf. ℓ_{jt} are random variables, as they depend on the probs, and in particular on $I_t, \dots, I_{t-1}$

We have $\bar{R}_T \leq \mathbb{E}[R_T]$.

We actually rather start for high-probability bounds on R_T , but studying $\bar{R}_T$ will be a good warm-up!



In these lecture notes, I'll take $N = K$ as the number of components

↳ we used N for individual sequences

↳ K stochastic bandits

and I alternatively took N and K in the next pages...

(My bad...)

Adversarial bandits:bound on $\bar{R}_T$ via exponential weights.

Key: Estimators of the losses (the unseen and the seen ones):

$$\hat{l}_{jt} = \frac{l_{j,t}}{p_{jt}} \mathbb{1}_{I_t=j} \quad \text{if } p_{jt} > 0 \quad \text{(which we will assume)}$$

auxiliary randomizations of opponent + decision-maker

They are (conditionally) unbiased: denoting by $\mathcal{F}_{t-1} = \sigma(U_1, \dots, U_{t-1}, l_{1,1}, \dots, l_{t-1,1}, p_{1,1}, \dots, p_{t-1,1}, I_1, \dots, I_{t-1})$

the total information available at the beginning of round t (of course, the decision-maker does not have that much information!), we have:

- l_t and p_t are $\mathcal{F}_{t-1}$ -measurable; the only randomness comes from the random draw of I_t according to p_t
- $\hat{l}_{jt}$ can be rewritten $\hat{l}_{jt} = \frac{l_{jt}}{p_{jt}} \mathbb{1}_{I_t=j}$

so that

$$\mathbb{E}[\hat{l}_{jt} | \mathcal{F}_{t-1}] = \frac{l_{jt}}{p_{jt}} \mathbb{E}[\mathbb{1}_{I_t=j} | \mathcal{F}_{t-1}] = \frac{l_{jt}}{p_{jt}} p_{jt} = l_{jt}$$

since we assumed $p_{jt} > 0$

Algorithm:

defined as

$p_1 = (1/N, \dots, 1/N)$ and for $t \geq 2$, $p_t = (p_{jt})_{j=1, \dots, N}$ is

for a non-increasing sequence $(\eta_t)_{t \geq 2}$

$$p_{jt} = \exp\left(-\eta_t \sum_{s=1}^{t-1} \hat{l}_{js}\right) / \sum_{k=1}^N \exp\left(-\eta_t \sum_{s=1}^{t-1} \hat{l}_{ks}\right)$$

L_1 ensures indeed that $p_{jt} > 0$.

the range $[0, M]$ is assumed to be known...

Theorem: The strategy above, tuned with $\eta_t = \frac{1}{M} \sqrt{\frac{\ln N}{Nt}}$, is such that:

for all opponents picking losses $l_{jt} \in [0, M]$,

$$\bar{R}_T = \mathbb{E}\left[\sum_{t=1}^T l_{I_t, t}\right] - \min_{i=1, \dots, N} \mathbb{E}\left[\sum_{t=1}^T l_{i, t}\right] \leq 2M \sqrt{T N \ln N}$$

The proof is based on the following lemma.

Lemma: The exponentially weighted average strategy on losses

$$\tilde{\ell}_j \in [0, +\infty[, \quad \text{i.e.,}$$

$$\text{with } \eta_t \downarrow, \quad \tilde{p}_{jt} = \exp\left(-\eta_t \sum_{s=1}^{t-1} \tilde{\ell}_{js}\right) / \sum_{k=1}^N \exp\left(-\eta_t \sum_{s=1}^{t-1} \tilde{\ell}_{ks}\right),$$

is such that

$$\sum_{t=1}^T \sum_{j=1}^N \tilde{p}_{jt} \tilde{\ell}_{jt} - \min_{i=1 \dots N} \sum_{t=1}^T \tilde{\ell}_{it} \leq \frac{\ln N}{\eta_T} + \sum_{t=1}^T \frac{\eta_t}{2} \sum_j \tilde{p}_{jt} \tilde{\ell}_{jt}^2.$$

Proof: We saw earlier in this series of lectures that the EWA strategy (with $\eta_t \downarrow$) is such that

$$\forall \tilde{\ell}_j \in \mathbb{R}, \quad \sum_{t=1}^T \tilde{p}_{jt} \tilde{\ell}_{jt} - \min_{i=1 \dots N} \sum_{t=1}^T \tilde{\ell}_{it} \leq \frac{\ln N}{\eta_T} + \sum_{t=1}^T \tilde{S}_t$$

$$\text{where } \tilde{S}_t = \sum_{j=1}^N \tilde{p}_{jt} \tilde{\ell}_{jt} + \frac{1}{\eta_t} \ln \sum_{j=1}^N \tilde{p}_{jt} e^{-\eta_t \tilde{\ell}_{jt}}$$

$$\text{We use here } e^{-x} \leq 1 - x + \frac{x^2}{2} \quad \forall x \geq 0$$

$$\begin{aligned} \text{so that } \ln \sum_j \tilde{p}_{jt} e^{-\eta_t \tilde{\ell}_{jt}} &\leq \ln \left(1 - \eta_t \sum_j \tilde{p}_{jt} \tilde{\ell}_{jt} + \frac{\eta_t^2}{2} \sum_j \tilde{p}_{jt} \tilde{\ell}_{jt}^2 \right) \\ &\stackrel{\substack{\ln(1+x) \leq x \\ \forall x > -1}}{\leq} - \eta_t \sum_j \tilde{p}_{jt} \tilde{\ell}_{jt} + \frac{\eta_t^2}{2} \sum_j \tilde{p}_{jt} \tilde{\ell}_{jt}^2 \end{aligned}$$

hence the stated bound.

Proof (of the theorem): We have no control on how large the $\tilde{\ell}_{jt}$ can be, and they could be very large! So, we would not be ready to apply any bound with a remainder $M_T \ln N$ term, where M_T is such that $\tilde{\ell}_{jt} \in [0, M_T]$ $\forall j, t$... as this M_T could be even super-linear. That's why we go back to the beginning of the

proof for the fully adaptive algorithm The η_t can be picked as $\ln N / \sum_{s=1}^{t-1} \delta_s$ or in terms of the upper bounds on the δ_s (we choose the latter version for the sake of concreteness).
 The lemma yields for the $\hat{L}_{jt}$: $\hookrightarrow$ see below!

$$\sum_{t,j} p_{jt} \hat{L}_{jt} - \min_{i=1,\dots,N} \sum_{t=1}^T \hat{L}_{it} \leq \frac{\ln N}{\eta_T} + \sum_{t=1}^T \frac{\eta_t}{2} \sum_j p_{jt} \hat{L}_{jt}^2$$

$\downarrow$ by definition of the $\hat{L}_{jt}$ $\downarrow$ similar treatment:

$$= \sum_j L_{jT} \frac{p_{jt}}{p_{jt}} \mathbb{1}_{jT=j} = L_{jT}$$

$$= \sum_j L_{jT}^2 \frac{1}{p_{jt}} \mathbb{1}_{jT=j} \leq M^2 \sum_j \mathbb{1}_{jT=j} / p_{jt}$$

To simplify even further the choice of the η_t , we first take $\mathbb{E}$ of both sides:

$$\mathbb{E} \left[\sum_t L_{jT} \right] - \underbrace{\mathbb{E} \left[\min_t \sum_t \hat{L}_{it} \right]}_{\leq \min_t \sum_t \mathbb{E}[\hat{L}_{it}] = \mathbb{E}[\hat{L}_{it}] \text{ by the tower rule and the fact that } \hat{L}_{it} \text{ is conditionally unbiased.}} \leq \frac{\ln N}{\eta_T} + \frac{M^2}{2} \sum_{t=1}^T \eta_t \underbrace{\sum_j \frac{\mathbb{E}[\mathbb{1}_{jT=j}]}{p_{jt}}}_{=1}$$

Thus,

$$\bar{R}_T = \mathbb{E} \left[\sum_{t=1}^T L_{jT} \right] - \min_{i=1,\dots,N} \mathbb{E} \left[\sum_{t=1}^T \hat{L}_{it} \right] \leq \frac{\ln N}{\eta_T} + \frac{M^2 N}{2} \sum_{t=1}^T \eta_t$$

the only adaptation to be made is just T (as M is assumed to be known)

The optimal constant η would be s.t. $\ln N / \eta = \frac{M^2 N}{2} T \eta$, that is,

$$\eta \text{ proportional to } \frac{1}{M} \sqrt{\frac{\ln N}{NT}}$$

$$\hookrightarrow \text{try } \eta_t = \frac{\gamma}{M} \sqrt{\frac{\ln N}{Nt}} \quad \text{where } \gamma \text{ is to be optimized.}$$

The bound is
$$M \sqrt{N \ln N} \left(\frac{\sqrt{T}}{\gamma} + \gamma \sum_{t=1}^T \frac{1}{\sqrt{t}} \right)$$

$$\leq \int_0^T \frac{1}{\sqrt{t}} dt \leq 2\sqrt{T}$$

$$\leq \sqrt{T} \left(\frac{1}{\gamma} + \gamma \right) = 2\sqrt{T} \quad \text{for } \gamma=1$$

Remarks / insights:

0 to apply
safely
 $e^{-x} \leq 1 - x + \frac{x^2}{2}$

M to bound $L_{\text{inf},t}^2 \leq M^2$
and to pick η_t

→ The range $[0, M]$ needs to be known; it could be a general range $[m, M]$, in which case, we would translate all losses by $-m$,
eg. consider $\hat{L}_{j,t} = \frac{L_{j,t} - m}{M - m} \mathbb{1}_{j \in J_t}$

and
$$\eta_t = \frac{1}{M - m} \sqrt{\frac{\ln N}{Nt}}$$

to get
$$R_T \leq 2(M - m) \sqrt{T N \ln N}$$

* Last-minute edit:
I think I have a solution
for R_T based on parts of
the additional document
↳ extra points for whoever finds it!

↳ Adaptation to the range in adversarial bandits is
actually a difficult issue (I would need to
think more about it...)*

→ In the next page you'll get a brief view of how to get high-probability bounds on
the true regret, of the same order of magnitude:

w.p. $1 - \delta$,
$$R_T \leq \square (M - m) \sqrt{T N \ln(N/\delta)}$$

→ There exists an algorithm (INF) s.t. R_T is controlled (w.h.p. or with high prob.) by $O(\sqrt{T N})$ against individual sequences $L_{j,t}$ [ie: sequences fixed in advance, not chosen by an opponent who reacts]

ie: no $\sqrt{\ln N}$ needed

↳ « oblivious » opponent

The $L_{j,t}$ are not random variables in this case.